# Computational Statistics and Data Analysis (MVComp2) <br> <br> Exercise 6 

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## 1 NYC taxicabs: frequentist inference (5 points)

While visiting New York city, you realize that each yellow taxicab displays a serial number. You assume that each cab $i$ displays a unique number, $x_{i}$, and that they are sequentially numbered starting from 1 . Can you infer the total number of taxicabs, $N$, given a set of $k$ observations, $x_{1}, x_{2}, \ldots, x_{k}$ ?
(a) Show that the conditional probability that the largest serial number observed is $M=m$, given that there are $N=n$ taxicabs and you make $K=k$ observations is given by

$$
P(M=m \mid N=n, K=k)=\left\{\begin{array}{rr}
\binom{m-1}{k-1}\binom{n}{k}^{-1}, & \text { if } k \leq m \text { and } m \leq n \\
0, & \text { otherwise }
\end{array}\right.
$$

(b) The expression in (a) is, in fact, the likelihood. Use maximum likelihood estimation to derive an estimator for $N$ as a function of $M$, denoted $\hat{N}_{1}(M)$. Is it a biased estimator?
(c) You propose to build a more robust estimator: Estimate the number of unobserved labels that are above the largest number observed, $M$. Assume that this number is equal to the average gap between observations. Show that your estimator for the total population size leads to

$$
\hat{N}_{2}(M)=\frac{k+1}{k} M-1
$$

(d) Use the likelihood in (a) to show that $\hat{N}_{2}$ is an unbiased estimator.
(e) The variance of the estimator is given by the expression

$$
\operatorname{Var}\left[\hat{N}_{2}\right]=\frac{1}{k} \frac{(n-k)(n+1)}{k+2} .
$$

In the regime of few observations, show that $\operatorname{Var}\left[\hat{N}_{2}\right]$ behaves in agreement to your assumptions.

## 2 NYC taxicabs: Bayesian inference (5 points)

Let's solve the same problem as in question 1, but using Bayesian inference. We want to use the likelihood in question 1 (a), together with an improper uniform prior over $N$, while fixing $K=k .{ }^{1}$
(a) Show that the posterior distribution, $P(N=n \mid M=m, K=k)$, is given by

$$
P(N=n \mid M=m, K=k)=\frac{k-1}{m}\binom{m}{k}\binom{n}{k}^{-1} .
$$

Hint: you may find the following Binomial coefficient identity useful

$$
\sum_{a=j}^{\infty}\binom{a}{b}^{-1}=\frac{b}{b-1} \frac{1}{\binom{j-1}{b-1}}
$$

(b) What is the maximum a-posteriori estimator?
(c) The posterior, $P(n \mid m, k)$, in fact corresponds to a shifted factorial distribution, such that $N-m \sim$ Fact $(k, m)$. A random variable, $Z$, follows a factorial distribution with parameters $n$ and $m$, i.e., $Z \sim$ Fact $(n, m)$, such that

$$
P(Z=z)=(n-1) \frac{(m-1)!}{(m-n)!} \frac{(m+z-n)!}{(m+z)!} .
$$

One can show that the expected value of $Z$ is given by $E[Z]=\frac{m-n+1}{n-2}$. Show that the posterior mean is given by

$$
\bar{N}=E[P(n \mid m, k)]=\frac{k-1}{k-2}(m-1) .
$$

(d) Consider the following sequence of serial numbers: $\boldsymbol{x}=(41,60,17,42)$. Compare the frequentist estimator, $\hat{N}_{2}$ in question 1 (c) with the present posterior mean, $\bar{N}$. Comment on the difference. What might be a more appropriate quantity for the posterior to better match the frequentist inference?

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[^0]:    ${ }^{1}$ An improper uniform prior is not bounded, and as such does not strictly speaking integrate to 1 over its domain. A more rigorous approach would consist of taking a large, but finite, interval. You can then show that the final result does not depend on the value of the bound. We will not do that here.

