Computational Statistics and Data Analysis (MVComp2)

Exercise 6

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1 NYC taxicabs: frequentist inference (5 points)

While visiting New York city, you realize that each yellow taxicab displays a serial number. You assume that each cab *i* displays a unique number, x_i , and that they are sequentially numbered starting from 1. Can you infer the total number of taxicabs, N, given a set of *k* observations, x_1, x_2, \ldots, x_k ?

(a) Show that the conditional probability that the largest serial number observed is M = m, given that there are N = n taxicabs and you make K = k observations is given by

$$P(M = m | N = n, K = k) = \begin{cases} \binom{m-1}{k-1} \binom{n}{k}^{-1}, & \text{if } k \le m \text{ and } m \le n \\ 0, & \text{otherwise} \end{cases}$$

- (b) The expression in (a) is, in fact, the likelihood. Use maximum likelihood estimation to derive an estimator for N as a function of M, denoted $\hat{N}_1(M)$. Is it a biased estimator?
- (c) You propose to build a more robust estimator: Estimate the number of unobserved labels that are *above* the largest number observed, *M*. Assume that this number is equal to the average gap between observations. Show that your estimator for the total population size leads to

$$\hat{N}_2(M) = \frac{k+1}{k}M - 1$$

- (d) Use the likelihood in (a) to show that \hat{N}_2 is an unbiased estimator.
- (e) The variance of the estimator is given by the expression

$$\mathrm{Var}[\hat{N}_2] = \frac{1}{k} \frac{(n-k)(n+1)}{k+2}.$$

In the regime of few observations, show that $\operatorname{Var}[\hat{N}_2]$ behaves in agreement to your assumptions.

2 NYC taxicabs: Bayesian inference (5 points)

Let's solve the same problem as in question 1, but using Bayesian inference. We want to use the likelihood in question 1 (a), together with an improper uniform prior over N, while fixing K = k.¹

(a) Show that the posterior distribution, P(N = n | M = m, K = k), is given by

$$P(N = n | M = m, K = k) = \frac{k - 1}{m} {\binom{m}{k}} {\binom{n}{k}}^{-1}.$$

Hint: you may find the following Binomial coefficient identity useful

$$\sum_{a=j}^{\infty} \binom{a}{b}^{-1} = \frac{b}{b-1} \frac{1}{\binom{j-1}{b-1}}.$$

- (b) What is the maximum a-posteriori estimator?
- (c) The posterior, P(n|m,k), in fact corresponds to a shifted factorial distribution, such that $N-m \sim Fact(k,m)$. A random variable, Z, follows a factorial distribution with parameters n and m, i.e., $Z \sim Fact(n,m)$, such that

$$P(Z=z) = (n-1)\frac{(m-1)!}{(m-n)!}\frac{(m+z-n)!}{(m+z)!}$$

One can show that the expected value of Z is given by $E[Z] = \frac{m-n+1}{n-2}$. Show that the posterior mean is given by

$$\bar{N} = E[P(n|m,k)] = \frac{k-1}{k-2}(m-1).$$

(d) Consider the following sequence of serial numbers: $\boldsymbol{x} = (41, 60, 17, 42)$. Compare the frequentist estimator, \hat{N}_2 in question 1 (c) with the present posterior mean, \bar{N} . Comment on the difference. What might be a more appropriate quantity for the posterior to better match the frequentist inference?

¹An *improper* uniform prior is not bounded, and as such does not strictly speaking integrate to 1 over its domain. A more rigorous approach would consist of taking a large, but finite, interval. You can then show that the final result does not depend on the value of the bound. We will not do that here.