# Computational Statistics and Data Analysis (MVComp2)

Solutions to exercise 3

 ${\color{black} \textbf{Lecturer}} \ {\color{black} {\rm Tristan \ Bereau}}$ 

Semester Wi23/24 Due Nov. 9, 2023, 23:59

### 1 Polluted water (2 points)

A particular concentration of a chemical found in polluted water has been found to be lethal to 20% of the fish that are exposed to the concentration for 24 hours. Twenty fish are placed in a tank containing this concentration of chemical in water.

- (a) Find the probability that at least 14 survive.
- (b) Find the mean and variance of the number that survive.

#### 1.1 Solution

First note that

$$P(X \ge 14) = 1 - P(X < 14) = 1 - P(X \le 13)$$

For  $P(X \le 13)$  use the binomial cumulative distribution function with survival probability 0.8:

$$P(X \le 13) = \sum_{k=0}^{13} \binom{20}{13} 0.8^{13} 0.2^{20-13}.$$

This yields a probability  $P(X \ge 14) \approx 91.3\%$ .

(b)

#### Mean

 $\mu = N\pi = 16$ 

Variance :

$$\mathrm{Var}[X] = N\pi(1-\pi) = 3.2$$

### 2 Distribution means (4 points)

(a) If X is a random variable with a geometric distribution  $P(X = x) = (1 - p)^{k-1}p$ , prove that the mean is given by

$$E[X] = \frac{1}{p}$$

(b) If X is a random variable with a Poisson distribution  $P(X = x) = \frac{\lambda^x}{x!} e^{-\lambda}$ , prove that the mean is given by

$$E[X] = \lambda$$

#### 2.1 Solution

(a)

$$E[X] = \sum_{y=1}^{\infty} yq^{y-1}p = p\frac{\mathrm{d}}{\mathrm{d}q}\left(\sum_{y=1}^{\infty} q^y\right) = p\frac{\mathrm{d}}{\mathrm{d}q}\left(\frac{q}{1-q}\right) = p\left[\frac{1}{(1-q)^2}\right] = \frac{1}{p}$$

(b)

$$E[X] = \sum_{i=0}^{\infty} i \frac{\lambda^i}{i!} e^{-\lambda} = \sum_{i=1}^{\infty} \frac{\lambda^i}{(i-1)!} e^{-\lambda} = \lambda \sum_{i=1}^{\infty} \frac{\lambda^{i-1}}{(i-1)!} e^{-\lambda} = \lambda \sum_{i=0}^{\infty} \frac{\lambda^l}{l!} e^{-\lambda} = \lambda$$

## 3 Conjugate priors (2 points)

Consider the Beta distribution with parameters  $\alpha$  and  $\beta$ 

$$Beta(\theta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{\alpha - 1} (1 - \theta)^{\beta - 1}$$

and the Binomial distribution for N observations

$$P(X = k) = \binom{N}{k} \theta^k (1 - \theta)^{N-k}$$

Show that the Beta distribution is a conjugate prior to the Binomial distribution. What are the parameters of the resulting distribution?

#### 3.1 Solution

Posterior from Bayes' rule:

$$\begin{split} p(\theta|X=k) &= p(X=k|\theta)p(\theta) \\ &= \binom{N}{k} \theta^k (1-\theta)^{N-k} \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{\alpha-1} (1-\theta)^{\beta-1} \\ &\propto \theta^{\alpha+k-1} (1-\theta)^{\beta+N-k-1}. \end{split}$$

The posterior indeed follows a Beta distribution, with parameters  $\alpha' = \alpha + k$  and  $\beta' = \beta + N - k$ .

### 4 The de Moivre–Laplace theorem (2 points)

The de Moivre–Laplace theorem states that the normal distribution may be used as an approximation to the Binomial distribution under certain conditions.<sup>1</sup> Let's find out!

Consider a random variable, X. The probability of getting k successes in n independent Bernoulli trials with probability p is given by the Binomial

$$P(X=k;n,p) = \binom{n}{k} p^k (1-p)^{n-k}.$$

The theorem states that the Binomial probability will converge to the normal probability density function,  $\mathcal{N}(np, \sqrt{np(1-p)})$ , as n becomes large and for p away from 0 or 1.

We will verify that the standardized random variable,  $\frac{X-np}{\sqrt{np(1-p)}}$ , approaches the standard normal,  $\mathcal{N}(0,1)$ , as n grows larger. Consider p = 0.5 and the cases  $n = \{2, 5, 10, 50, 100\}$ . Plot all curves on the same graph to show that they overlap.

#### 4.1 Solution

```
import numpy as np
from scipy.stats import binom, norm
import matplotlib.pyplot as plt
fig, ax = plt.subplots(1, 1)
plt.xlim([-4., 4.])
x_norm = np.linspace(*plt.xlim(), num=100)
ax.plot(
  x_norm, norm.pdf(x_norm, loc=0., scale=1.), "--", label="unit Gaussian"
)
p = 0.50
ns = [2, 5, 10, 50, 100]
for n in ns:
  mean = n * p
  std = np.sqrt(n * p * (1.-p))
  # Generate a range for x (from 0 to n) and calculate the binomial probability
  x = np.arange(0, n+1)
  y_binom = binom.pmf(x, n, p)
  # Standardize the x values (convert to z-scores)
  z_scores = (x - mean) / std
  # Convert probability to a PDF-like form by scaling with std
```

<sup>&</sup>lt;sup>1</sup>This is a special case of the central limit theorem. A famous physical realization of the de Moivre–Laplace theorem is the Galton box, also called bean machine.

```
y_scaled = y_binom * std
  # Plot the binomial distributions as step functions
  ax.plot(z_scores, y_scaled, drawstyle='steps-mid', label=f'Binomial n={n}')
plt.grid()
plt.legend()
plt.xlabel(r"$z$")
```

```
plt.ylabel(r"$P(z)$")
plt.show()
```



