# Computational Statistics and Data Analysis (MVComp2) 

## Solutions to exercise 3

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## 1 Polluted water (2 points)

A particular concentration of a chemical found in polluted water has been found to be lethal to $20 \%$ of the fish that are exposed to the concentration for 24 hours. Twenty fish are placed in a tank containing this concentration of chemical in water.
(a) Find the probability that at least 14 survive.
(b) Find the mean and variance of the number that survive.

### 1.1 Solution

(a)

First note that

$$
P(X \geq 14)=1-P(X<14)=1-P(X \leq 13)
$$

For $P(X \leq 13)$ use the binomial cumulative distribution function with survival probability 0.8 :

$$
P(X \leq 13)=\sum_{k=0}^{13}\binom{20}{13} 0.8^{13} 0.2^{20-13}
$$

This yields a probability $P(X \geq 14) \approx 91.3 \%$.
(b)

## Mean

$$
\mu=N \pi=16
$$

Variance :

$$
\operatorname{Var}[X]=N \pi(1-\pi)=3.2
$$

## 2 Distribution means (4 points)

(a) If $X$ is a random variable with a geometric distribution $P(X=x)=(1-p)^{k-1} p$, prove that the mean is given by

$$
E[X]=\frac{1}{p}
$$

(b) If $X$ is a random variable with a Poisson distribution $P(X=x)=\frac{\lambda^{x}}{x!} \mathrm{e}^{-\lambda}$, prove that the mean is given by

$$
E[X]=\lambda
$$

### 2.1 Solution

(a)

$$
E[X]=\sum_{y=1}^{\infty} y q^{y-1} p=p \frac{\mathrm{~d}}{\mathrm{~d} q}\left(\sum_{y=1}^{\infty} q^{y}\right)=p \frac{\mathrm{~d}}{\mathrm{~d} q}\left(\frac{q}{1-q}\right)=p\left[\frac{1}{(1-q)^{2}}\right]=\frac{1}{p}
$$

(b)

$$
E[X]=\sum_{i=0}^{\infty} i \frac{\lambda^{i}}{i!} \mathrm{e}^{-\lambda}=\sum_{i=1}^{\infty} \frac{\lambda^{i}}{(i-1)!} \mathrm{e}^{-\lambda}=\lambda \sum_{i=1}^{\infty} \frac{\lambda^{i-1}}{(i-1)!} \mathrm{e}^{-\lambda}=\lambda \sum_{i=0}^{\infty} \frac{\lambda^{l}}{l!} \mathrm{e}^{-\lambda}=\lambda
$$

## 3 Conjugate priors (2 points)

Consider the Beta distribution with parameters $\alpha$ and $\beta$

$$
\operatorname{Beta}(\theta)=\frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha) \Gamma(\beta)} \theta^{\alpha-1}(1-\theta)^{\beta-1}
$$

and the Binomial distribution for $N$ observations

$$
P(X=k)=\binom{N}{k} \theta^{k}(1-\theta)^{N-k}
$$

Show that the Beta distribution is a conjugate prior to the Binomial distribution. What are the parameters of the resulting distribution?

### 3.1 Solution

Posterior from Bayes' rule:

$$
\begin{aligned}
p(\theta \mid X=k) & =p(X=k \mid \theta) p(\theta) \\
& =\binom{N}{k} \theta^{k}(1-\theta)^{N-k} \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha) \Gamma(\beta)} \theta^{\alpha-1}(1-\theta)^{\beta-1} \\
& \propto \theta^{\alpha+k-1}(1-\theta)^{\beta+N-k-1} .
\end{aligned}
$$

The posterior indeed follows a Beta distribution, with parameters $\alpha^{\prime}=\alpha+k$ and $\beta^{\prime}=\beta+N-k$.

## 4 The de Moivre-Laplace theorem (2 points)

The de Moivre-Laplace theorem states that the normal distribution may be used as an approximation to the Binomial distribution under certain conditions. ${ }^{1}$ Let's find out!

Consider a random variable, $X$. The probability of getting $k$ successes in $n$ independent Bernoulli trials with probability $p$ is given by the Binomial

$$
P(X=k ; n, p)=\binom{n}{k} p^{k}(1-p)^{n-k} .
$$

The thoerem states that the Binomial probability will converge to the normal probability density function, $\mathcal{N}(n p, \sqrt{n p(1-p)})$, as $n$ becomes large and for $p$ away from 0 or 1 .
We will verify that the standardized random variable, $\frac{X-n p}{\sqrt{n p(1-p)}}$, approaches the standard normal, $\mathcal{N}(0,1)$, as $n$ grows larger. Consider $p=0.5$ and the cases $n=\{2,5,10,50,100\}$. Plot all curves on the same graph to show that they overlap.

### 4.1 Solution

```
import numpy as np
from scipy.stats import binom, norm
import matplotlib.pyplot as plt
fig, ax = plt.subplots(1, 1)
plt.xlim([-4., 4.])
x_norm = np.linspace(*plt.xlim(), num=100)
ax.plot(
    x_norm, norm.pdf(x_norm, loc=0., scale=1.), "--", label="unit Gaussian"
)
p = 0.50
ns = [2, 5, 10, 50, 100]
for n in ns:
    mean = n * p
    std = np.sqrt(n * p * (1.-p))
    # Generate a range for x (from O to n) and calculate the binomial probability
    x = np.arange(0, n+1)
    y_binom = binom.pmf(x, n, p)
    # Standardize the x values (convert to z-scores)
    z_scores = (x - mean) / std
    # Convert probability to a PDF-like form by scaling with std
```

[^0]y_scaled = y_binom * std
\# Plot the binomial distributions as step functions
ax.plot(z_scores, y_scaled, drawstyle='steps-mid', label=f'Binomial n=\{n\}')

```
plt.grid()
plt.legend()
plt.xlabel(r"$z$")
plt.ylabel(r"$P(z)$")
plt.show()
```




[^0]:    ${ }^{1}$ This is a special case of the central limit thoerem. A famous physical realization of the de Moivre-Laplace theorem is the Galton box, also called bean machine.

